

Adaptive design of cluster randomised trials

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Outline

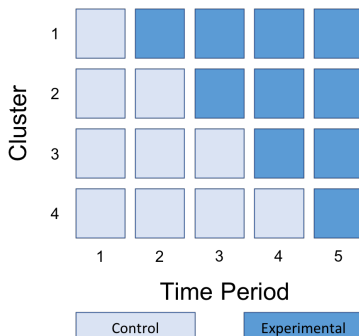
- 1 Introduction
- 2 Design framework
- 3 Group sequential SW-CRTs
- 4 Sample size re-estimation in SW-CRTs
- 5 Discussion

Adaptive design of PG-CRTs

- Much research has now been conducted to facilitate adaptive trial design
- But the majority of focus has been on individual level randomisation
- Some exceptions for parallel group (PG) cluster randomised trials (CRTs)
 - Lake *et al* (2002) discussed re-estimation of the required number of clusters
 - Zou *et al* (2005) described group sequential design for binary outcomes
 - van Schie and Moerbeek (2014) considered re-estimation of the required number of individuals per cluster

CRXO and SW-CRTs

- In recent years there has been increased interest in cluster randomised crossover (CRXO) and stepped-wedge (SW) CRTs



- Motivation comes from considerations related to these designs

CRXO and SW-CRTs

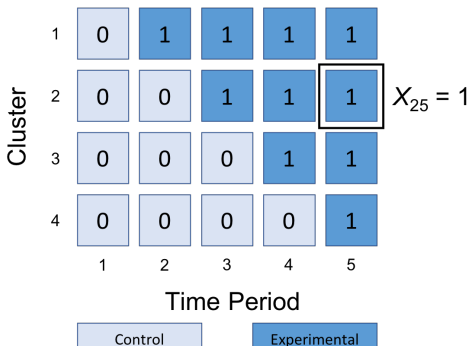
- 31% of SW-CRTs failed to find efficacy on their primary outcome
- Possibly associated with over enthusiastic use of the design
- Or could be due to the challenge of specifying the variance components
- Describe a flexible framework for incorporating interim assessments of futility/efficacy in SW-CRTs/CRXO trials
- Detail how we can re-estimate the required sample size in a blinded or unblinded manner

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Trial setting

- Consider a SW-CRT in C clusters over T time periods, with m measurements per cluster per period
- Suppose the treatment allocations have also been specified, via a matrix X of zeroes and ones



Analysis

- Can be used when there are different participants in each period, or the same participants
- They also work for almost any choice of linear mixed model
- For simplicity

$$Y_{ijk} = \mu + \pi_j + \tau X_{ij} + c_i + e_{ijk}.$$

- The random cluster effect $c_i \sim N(0, \sigma_c^2)$, and the residual error $e_{ijk} \sim N(0, \sigma_e^2)$
- We are interested in testing

$$H_0 : \tau \leq 0, \quad H_1 : \tau > 0.$$

- We want a type-I error-rate of α when $\tau = 0$, and power of $1 - \beta$ when $\tau = \delta$

Example

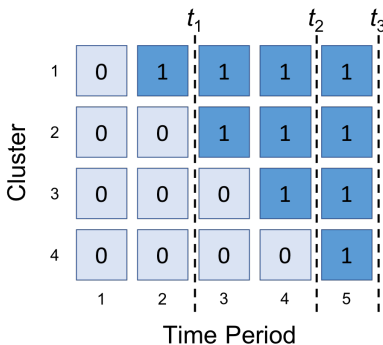
- Bashour *et al* (2013) described a SW-CRT to assess the effect of training doctors in communication skills on women's satisfaction with doctor-woman relationship during labour and delivery
- Design had $C = 4$, $T = 5$, estimated $\hat{\tau} = -0.13$, $\hat{\sigma}_c^2 = 0.02$ and $\hat{\sigma}_e^2 = 0.51$
- For $\alpha = 0.05$, $\beta = 0.1$, $\delta = 0.2$, design required $m = 70$ patients per cluster per period
- Use these as the example parameters in all of what follows

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Goal

- Include interim analyses where we can test for futility/efficacy, reducing the required number of observations
- Specify a collection of time periods after which we will conduct analyses: t_1, t_2, \dots



Stopping rules

- Our testing rules are then based on efficacy and futility stopping boundaries, $e = (e_{t_1}, e_{t_2}, \dots)$ and $f = (f_{t_1}, f_{t_2}, \dots)$
- We need a test statistic Z_t to use after time period t
 - If $Z_t > e_t$ then we stop the trial, and reject H_0
 - If $Z_t \leq f_t$ then we stop the trial, and do not reject H_0
 - Otherwise we continue the trial to period $t + 1$
- Natural to use a Wald test statistic

$$Z_t = \frac{\hat{\tau}_t}{\sqrt{\text{Var}(\hat{\tau}_t)}} = \hat{\tau}_t I_t^{1/2}.$$

Theory

- Established group sequential design methodology is directly applicable to this longitudinal setting
- $\mathbf{Z} = (Z_{t_1}, Z_{t_2}, \dots)$ has a multivariate normal distribution

$$\mathbb{E}(Z_t) = \tau I_t^{1/2},$$

$$\text{Cov}(Z_{t_i}, Z_{t_j}) = (I_{t_i}/I_{t_j})^{1/2}, \quad t_i \leq t_j.$$

- So we can compute the probability we stop for efficacy/futility at each analysis using multivariate normal integration
- Adding up the probability you stop for efficacy at each analysis gives you the overall rejection probability
- Use this in...

Design determination

- We need to choose m , e , and f , given choices of the variance components σ_c^2 and σ_e^2 , that provide the desired error-rates
- A simple solution is to use the “error spending” approach to sequential trial design
- Or a global optimisation algorithm can be used to find the best possible boundaries

- Created some simple software for this

```
> example <- gs_sw(X, sigma_e, sigma_c, alpha,
                  beta, delta, set_T)
```

```
> summary(example)
```

Identified design has:

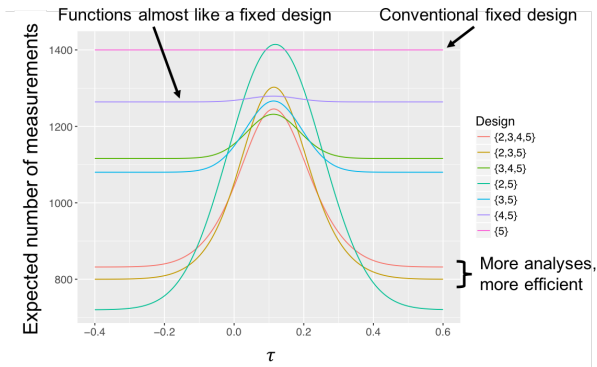
$m = 97,$

$e = (2.32, 2.06, 1.90),$

$f = (-0.57, 0.65, 1.90).$

Expected number of measurements

- Consider the influence of the choice of analysis times



- But** including interim analyses increases the maximal possible required number of measurements

Summary

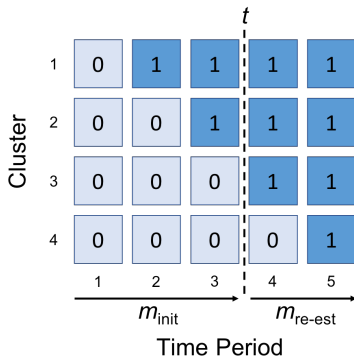
- Long-standing group sequential design theory can be used to incorporate interim analyses in to SW-CRTs
- Could be an effective method for reducing the expense of such trials when an intervention is highly effective/ineffective
- To power correctly, requires an assumption of known variance...

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Goal

- After some specified time period t , re-estimate the required variance components
- Go from m_{init} measurements per cluster per period to a hopefully more appropriate $m_{\text{re-est}}$



Unblinded re-estimation

- Always straight forward to implement, but less favoured by regulatory agencies
- Fit the chosen model to the data to acquire the re-estimates $\hat{\sigma}_c^2$ and $\hat{\sigma}_e^2$
- Can write down a variance for the information when m_{init} patients are used in time periods $1, \dots, t$, and $m_{\text{re-est}}$ in time periods $t + 1, \dots, T$

$$\text{Var}(\hat{\tau} \mid m_{\text{init}}, m_{\text{re-est}}, \hat{\sigma}_c^2, \hat{\sigma}_e^2).$$

- Use these in the conventional method of sample size determination

$$\Phi\{\delta/\text{Var}(\hat{\tau} \mid m_{\text{init}}, m_{\text{re-est}}, \hat{\sigma}_c^2, \hat{\sigma}_e^2) - z_{1-\alpha}\} \geq 1 - \beta.$$

Blinded re-estimation

- Works in a similar way, but to acquire out estimates we use

$$S_{\text{eq1}}^2 = \frac{m_{\text{init}}}{Ct - t} \sum_{i=1}^C \sum_{j=1}^t (\bar{Y}_{ij.} - \bar{Y}_{.j.})^2,$$

$$S_{\text{eq2}}^2 = \frac{1}{m_{\text{init}}Ct - Ct} \sum_{i=1}^C \sum_{j=1}^t \sum_{k=1}^{m_{\text{init}}} (Y_{ijk} - \bar{Y}_{ij.})^2.$$

- We can show that in the absence of a treatment effect

$$\hat{\sigma}_e^2 = S_{\text{eq2}}^2,$$

$$\hat{\sigma}_c^2 = \frac{1}{n} (S_{\text{eq1}}^2 - S_{\text{eq2}}^2),$$

are unbiased estimates of the variance components

Simulation study: Power

Assumed variances	Blinded	Unblinded	Fixed
50% smaller than truth	0.881	0.880	0.692
Correct	0.885	0.884	0.903
50% larger than truth	0.897	0.895	0.974

Procedure	Assumed variances	$t = 2$	$t = 3$	$t = 4$
Blinded	50% smaller than truth	0.870	0.881	0.815
Blinded	Correct	0.885	0.885	0.897
Blinded	50% larger than truth	0.888	0.897	0.954
Unblinded	50% smaller than truth	0.869	0.880	0.817
Unblinded	Correct	0.884	0.884	0.897
Unblinded	50% larger than truth	0.887	0.896	0.955

Summary

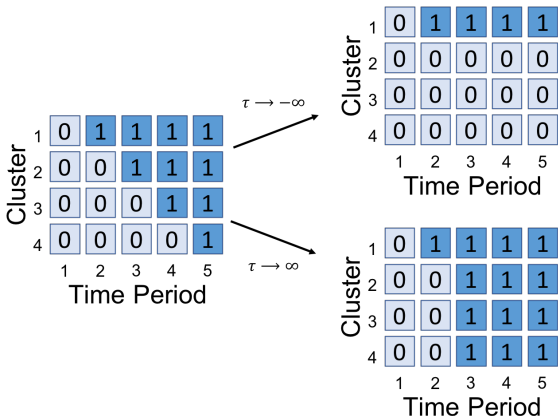
- Sample size re-estimation can greatly increase your power when you have under-specified the variance components
- Only a single re-estimation point required. Feasible in practice?
- Need to think carefully about when to time the re-estimation
- One other issue is that it is quite computationally intensive to investigate these designs

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Discussion

- A variety of adaptive designs are available for CRTs
- Also working on a technique for response adaptive treatment allocation



Discussion

- Time period structure of SW-CRTs may make interim analyses particularly appealing
- However, there are issues associated with their use
- Can interim analyses be handled efficiently?
- If you stop for efficacy, what would you do next if resources are scarce?
- Can you realistically handle an increase in the cluster period sample size?
- Nonetheless, if used wisely they could help greatly with improving efficiency/power of SW-CRTs

References

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